

Are Psychophysical Scales of Intensities the Same or Different When Stimuli Vary on Other Dimensions? Theory With Experiments Varying Loudness and Pitch

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Most studies concerning psychological measurement scales of intensive attributes have concluded that these scales are of ratio type and that the psychophysical function is closely approximated by a power function. Experiments show, for such cases, that a commutativity property must hold under either successive increases or successive decreases provided, e.g., all other independent dimensions are fixed. A good deal of data support this conclusion. However, little or no attention has been paid to whether or not such subjective intensity scales differ when an independent dimension such as frequency (pitch in audition, color in vision, etc.) is varied. Using a simple and favorably tested theoretical model for global psychophysics, the authors arrive at a necessary and sufficient cross-dimension, commutativity condition for a common intensity ratio scale to exist. For example, the data show that the loudness of a tone at frequency f and another tone at frequency g can each be viewed as arising from a common property of loudness over intensity/frequency pairs. Comparing one version of cross-dimensional commutativity with the corresponding 1-dimensional commutativity property discriminates between a general representation of the ratio scale property and a special case of it. Future work: Does the theory extend to other intensive continua (prothetic attributes)? If so, which ones? And does it extend to cross-modal matching?

Keywords: pitch, frequency, cross modal matching, loudness, scaling

S. S. Stevens (1946, 1951) produced a new theory of measurement that allowed for the direct measurement of subjective intensities using the now well-known methods of magnitude estimation and production. These methods have been and continue to be widely used by psychologists and other behavioral scientists. Yet Stevens's theory and his methods of measurement by magnitude estimation and production are considered by the current standards of measurement theorists, both within and outside of psychology, and by philosophers of science and others interested in the theory of measurements to be nonrigorous and/or ill-founded.

Narens (1996) offered an axiomatic analysis of Stevens's methods using the representational theory of measurement, and he discovered alternatives to Stevens's methods, which he showed were subjectively measured rigorously on ratio scales. The principal difference between two forms of measurement is that Stevens assumed that the respondents give a veridical interpretation of the instructions to produce either ratio estima-

tions or ratio productions, whereas Narens required them to be represented as some ratio but not necessarily the exact ratio requested in the instructions. Later, other theories (Augustin, 2006; Luce, 2004, 2008; Narens, 1997, 2006) expanded the Narens (1996) approach. Further, the experimental results of several researchers cited throughout this article have established the predicted deficiencies of Stevens' approach and supported the predicted strengths of both Luce's and Narens's theoretical approaches.

Luce's and Narens's approaches are examples of a theory of subjective intensity called the *relation theory*. The relation theory (Krantz, 1972; Shepard, 1981) asserts that the fundamental object judged by an observer is a pair of stimuli, rather than a single stimulus. We apply the relation theory to pairs of stimuli from different dimensions in a manner that is different from others in the literature (e.g., Krantz, 1972).

Narens's models rest on the idea that ratio scalability is essentially equivalent to a simple commutativity property for signals varying only on a single dimension of intensity. For loudness magnitude production, the commutativity property is formulated as follows: Let x be an arbitrary sound intensity less its threshold intensity, and let p and q be two positive numbers. The respondent is asked to produce a sound intensity x_p that is heard as " p times as loud as x " and later to produce a sound intensity $x_{p,q}$ that is " q times as loud as x_p ." Next suppose that the respondent is asked to reverse the order so that q is first and p second, i.e., to produce a sound x_q that is " q times as loud as x " and then a sound $x_{q,p}$ that is " p times as loud as x_q ." The commutativity property holds if and only if $x_{p,q} = x_{q,p}$. As we said, this prediction has been supported empirically in several studies (see the following).

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Luce's (2004) global psychophysical model is somewhat richer in concepts than is Narens's model, and so it applies to a richer set of experimental situations. Also, it has been extensively and favorably evaluated experimentally in loudness and brightness (Luce & Steingrimsson, 2008; Steingrimsson, 2009, in press; Steingrimsson & Luce, 2005a, 2005b, 2006, 2007). Further, Narens's model, when viewed as a model of the psychophysical function, becomes a special case of Luce's where the reference points are set to 0. Moreover, Luce's and Narens's models both predict commutativity, but the former appears to be more straightforward to describe and has somewhat greater explanatory powers. So we have chosen it as the basis of the theory presented in this article and have used it to account for the experimental data we report.

Narens (2006) formulated various ways to extend the commutativity property and used those extensions to conclude when it is possible to represent various ratio scales all on a common ratio scale. The latter is a general measurement-theoretic result that also applies outside of psychophysical situations. To apply the underlying idea to magnitude production when the ratio scales are on disjoint domains, we needed to develop a new kind of paradigm. We did so using a theoretically motivated paradigm that can be applied to a wide variety of psychophysical situations, including ones involving mixed modalities. The experimental part of this article demonstrates the effectiveness of the new paradigm by investigating whether or not frequency-dependent ratio scales for auditory loudness can be represented on a common ratio scale.

The paradigm can be used equally well to address more general questions. For example, assuming that brightness turns out to have properties similar to the loudness ones discussed in this article, can loudness and brightness each be considered to be special cases of a common ratio scale of subjective magnitude for both modalities? If so, to what other domains does cross-dimensional commutativity generalize? The most general case would be to have a common scale for all intensity magnitudes classed as prothetic by Stevens (1975). This could develop into quite an extensive experimental exploration.

The article begins with a concise statement of the theory, describes the experimental design, and reports the results and their implications.

A Psychophysical Theory of Ratio Scales

From several behavioral (testable) axioms (i.e., putative laws), Luce (2004) showed that when a respondent is asked to produce the signal x_p that stands in ratio p to a given signal x , then there exist two functions: ψ , a strictly increasing, psychophysical ratio scale over the set of signal intensities, and W , a cognitive distortion over numbers, as well as a parameter ρ_i , called a reference signal (discussed below), that satisfies the constraint

$$W(p) = \frac{\psi(x_p) - \psi(\rho_i)}{\psi(x) - \psi(\rho_i)}, \quad i = \begin{cases} + & \text{if } p \geq 1 \\ - & \text{if } p < 1 \end{cases} \quad (1)$$

For our purposes, we treat the reference signal ρ_i as a physical signal generated by the respondent; these can, with suitable sets of data, be estimated as detailed later. The reason for assuming that $\rho_+ \neq \rho_-$ is that the data themselves require it. It would be most desirable to have a theory for the reference signals, but currently we do not have one.

Our theoretical approach is modeled on classical physical measurement, which completely ignores the internal forces among atoms and attempts only to describe the macro behavior of physical measures, such as mass, density, and force. So we make no attempt to relate Representation 1 to any mind or brain activities.

Throughout the article, we examine successive productions. For example, with x_p satisfying Representation 1, then $x_{p,q}$ is constructed relative to x_p by

$$W(q) = \frac{\psi(x_{p,q}) - \psi(\rho_i)}{\psi(x_p) - \psi(\rho_i)}, \quad i = \begin{cases} + & \text{if } q \geq 1 \\ - & \text{if } q < 1 \end{cases} \quad (2)$$

The focus of the article is on whether or not commutativity obtains in the sense $x_{p,q} = x_{q,p}$. We show in Proposition 1 that this indeed follows from Representation 1 and Equation 2.

Representation 1 was derived from axioms including certain structural ones, such as solvability, and others that are behavioral and experimentally testable. One formulation of these conditions is given as Corollary 1 of Theorem 1 of Luce (2004). These testable properties have been evaluated favorably for both loudness and brightness (study of perceived contrast is underway). For such a simple representation, other axiomatic formulations almost certainly will be discovered in the future.

For some purposes, it is convenient to rewrite Representation 1 in a linear form:

$$\psi(x_p) = W(p)\psi(x) + [1 - W(p)]\psi(\rho_i). \quad (3)$$

Those familiar with the large literature on utility representations of risky decision making will recognize Equation 3 as subjective expected utility for binary lotteries, provided that $0 \leq p \leq 1$ and $0 \leq W(p) \leq 1$.

Further, as verified empirically by Steingrimsson and Luce (2005a, 2005b), Luce & Steingrimsson (2008), and Steingrimsson (2009, in press), when a respondent selects x_p to be the signal that is perceived to be p times that of x relative to its reference intensity ρ (magnitude production), the data are nicely described by this representation. Having an estimate of x_p , it then can play the role of x under the instruction to find $x_{p,q}$ that is q times x_p . Such data permit one to examine the commutativity property.

Steingrimsson and Luce (2006) provided data supporting the idea that ψ may be a power function, i.e.,

$$\psi(x) = \alpha x^\beta, \quad \alpha > 0, \beta > 0. \quad (4)$$

We very much need to devise a means for estimating β for each respondent. We do not need to estimate α because it cancels in Representation 1.

Steingrimsson and Luce (2007) explored the form for W and supported the representation of Prelec (1998) but used a simpler property because of Luce (2001). We do not report this result in detail because it plays no role in this article.

Such restrictions on ψ and W reduced the degrees of freedom in Representation 1 to six free parameters.

Two Propositions for Single Frequency Case

Because we are using two numbers p and q , we need to be careful to distinguish four cases: both $p \geq 1$ and $q \geq 1$; both $p' < 1$ and $q' < 1$; and the two mixed cases, $p \geq 1$ and $q' < 1$ and $p' < 1$ and $q \geq 1$. Note the we prime any number less than 1.

Sometimes (see Representation 1 and Case 3 of Proposition 1) we use $+$ and $-$ to refer to the sign of $p - 1$.

Our first two Propositions concern the single intensity dimension shown as Case 1 in Figure 1.

Proposition 1. Assuming that the general model (Representation 1) holds, then in the one-dimensional case,

1. For $p \geq 1$ and $q \geq 1$, commutativity $x_{p,q} = x_{q,p}$ holds.
2. For $p' < 1$ and $q' < 1$, commutativity $x_{p',q'} = x_{q',p'}$ holds.
3. For $p' < 1 \leq q$ and for $p \geq 1 > q'$, commutativity holds iff $\rho_+ = \rho_-$.

Proofs are given in Appendix A.

Cases 1 and 2 have been extensively studied and supported (Augustin & Maier, 2008; Ellermeier & Faulhammer, 2000; Steingrímsson, 2009; Steingrímsson & Luce, 2005a; Zimmer, 2005), so we take Cases 1 and 2 to be confirmed.

Case 3 has not been systematically studied empirically. Nothing that we know of precludes that $\rho_+ = \rho_-$, but equally well, nothing a priori forces it. There has been very little study of the mixed case except for a limited amount of data reported in Appendix E of Steingrímsson and Luce (2007) that imply that $\rho_+ \neq \rho_-$.

The next proposition provides a means for estimating ρ_+ and ρ_- . Define

$$\tau_+(p, q) := W(p)W(q) \text{ if } p > 1, q > 1, \quad (5)$$

$$\tau_-(p', q') := W(p')W(q') \text{ if } p' < 1, q' < 1. \quad (6)$$

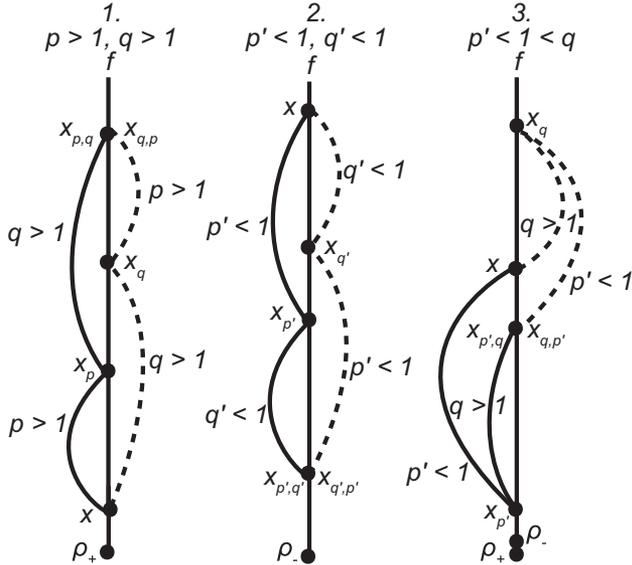


Figure 1. Depicted are the three cases of Proposition 1. In each graph, the solid lines depict the case of magnitude production with p followed by q , and the dotted ones depict the case of magnitude production with q followed by p . Commutativity is said to hold in Case 1 if $x_{p,q} \sim x_{q,p}$, in Case 2 if $x_{p',q'} \sim x_{q',p'}$, and in Case 3 if $x_{p',q} \sim x_{q,p'}$. The reference point for $p, q > 1$ is ρ_+ and for $p', q' < 1$, it is ρ_- . In Case 3, both reference points are in play, but commutativity is predicted only if $\rho_+ = \rho_-$.

Proposition 2. Under the assumption of Proposition 1 and assuming Equation 4, the parameters of Cases 1 and 2 must satisfy for all intensities x and empirically estimated $x_{p,q}$:

$$\rho_+^\beta = \frac{x^\beta \tau_+(p, q) - x_{p,q}^\beta}{\tau_+(p, q) - 1}, \quad (7)$$

$$\rho_-^\beta = \frac{x^\beta \tau_-(p', q') - x_{p',q'}^\beta}{\tau_-(p', q') - 1}. \quad (8)$$

With p, q (and p', q') held fixed, the data as we vary x should satisfy the linear expression (Equation 3), and so we may estimate ρ_i and τ_i separately for $i = +, -$. The separateness rests upon the empirical fact that respondents do behave differently. Because estimating beta is outside the scope of this article, only the means of estimating the reference points is presented. These are not estimated in this article.

Cross-Frequency Commutativity

General notation. The above arguments are readily modified to two frequency dimensions, say f and g , as in Figure 2, by suitable super- and subscripting. For example, in the single frequency case, we wrote x_p , whereas had we been explicit that we were mapping f to f , we should have written $x_p^{f,f}$. In the case where the judgment is from frequency f to frequency g , that must be explicit as $x_p^{f,g}$. Then, as long as the roles of p and q are simply switched, the proposition continues to be satisfied.

The notation x^f simply indicates that x has the frequency f , not a power of x . The same is true of $x_p^{f,g}$. Because we are restricting ourselves to the two frequencies f and g , this convention appears the least cumbersome of those we have considered.¹

More specifically, assume that loudness at each frequency is correctly described by Luce's (2004, 2008) theory. Let the resulting numerical ratio scales of loudness at f and g be denoted, respectively, by ψ^f and ψ^g (again, f and g refer to frequencies, not powers). Recall that each scale is assumed to have a reference point, ρ^f and ρ^g , from which the judgments are made.

We assume that the claim that ψ^f and ψ^g are the same ratio scale means that there is a single ratio scale $\psi^{f,g}$ with arguments (x, h) for intensity x at frequency h that agrees with ψ^f when $h = f$ and with ψ^g when $h = g$.

Suppose a signal intensity x at frequency f is presented and the respondent reports the magnitude production intensity at frequency g that is perceived as p times x . We denote that production by $x_p^{f,g}$. If the next step is to produce on g the signal that is $q \times x_p^{f,g}$, the resulting signal is denoted $x_{p,q}^{f,g}$. If, however, the second magnitude production is at frequency f , the notation changes to $x_{p,q}^{f,g,f}$. When the roles of p and q are reversed, the notation is appropriately changed to reflect this.

We assume that the reference point chosen depends both on whether or not $p \geq 1$ or $p' < 1$ and on whether it concerns the f or g frequency.

In the following, we write $f \rightarrow g$ to mean we are considering a cross-frequency production from frequency f to frequency g .

¹ We use the simple f, g notation. However, were we to consider a property with $n > 2$ different dimensions, we would almost certainly switch to writing $f_j, j = 1, \dots, n$. For $n = 2$, we avoid the additional subscripts.

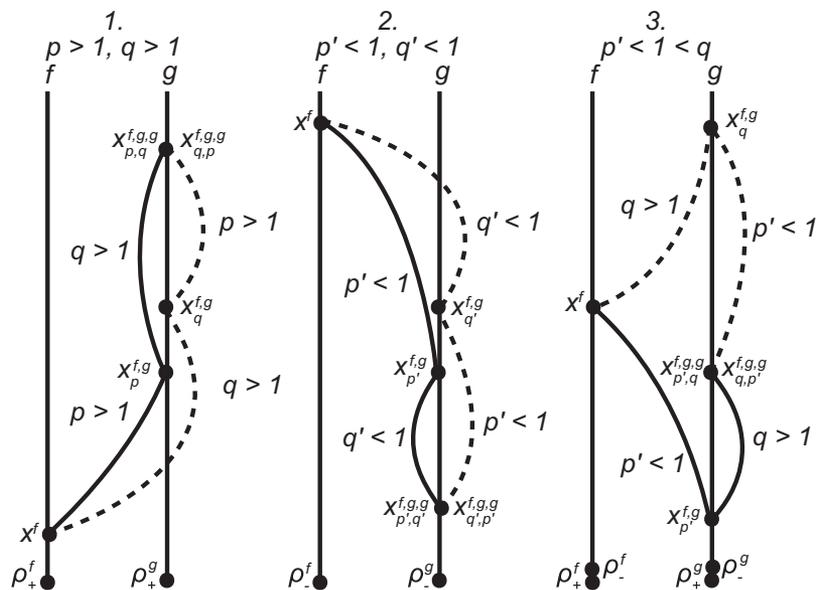


Figure 2. Depicted are the three cases of Proposition 1 when extended to two frequencies, f , g . In each of the three cases, the solid lines depict the case of magnitude production with p followed by q , and the dotted lines depict magnitude production starting with q followed by p . Commutativity is found to hold in Case 1 if $x_{p,q}^{f,g,g} \sim x_{q,p}^{f,g,g}$, in Case 2 if $x_{p',q'}^{f,g,g} \sim x_{q',p'}^{f,g,g}$, and in Case 3 if $x_{p',q}^{f,g,g} \sim x_{q,p'}^{f,g,g}$. The reference point for p , $q > 1$ is ρ_+ , and for p' , $q' < 1$, it is ρ_- . In Case 3, both reference points are in play, but commutativity is predicted only if $\rho_+ = \rho_-$.

The cases $f \rightarrow g \rightarrow g$ and $f \rightarrow g \rightarrow f$. Three distinct cross-frequency cases, called 1, 2, and 3, are diagrammed in Figure 2. Note that in these cases, there is one cross-dimension magnitude production followed by single-dimension magnitude production. Figure 2 presents the realizations that we found to be most effective (it parallels Figure 1 for testing of single frequency); the first item we study begins with a judgment in frequency g of the given intensity at frequency f , namely $x_p^{f,g}$.

Proposition 3. Assuming that the general model (Representation 1) holds, the cross mapping $f \rightarrow g \rightarrow g$ and $f \rightarrow g \rightarrow f$ both satisfy cross-dimensional commutativity for Cases 1 and 2 of Proposition 1 and hold for Case 3 if and only if $\rho_+^f = \rho_-^f$ and $\rho_+^g = \rho_-^g$.

There is an additional case to explore.

Comparison of $f \rightarrow f \rightarrow f$ with $f \rightarrow g \rightarrow f$. We have already shown commutativity on dimension f ,

$$x_{p,q} = x_{q,p}$$

and commutativity in the two-dimensional case $f \rightarrow g \rightarrow f$,

$$x_{p,q}^{f,g,f} = x_{q,p}^{f,g,f}$$

A natural question to ask is, “Do these two cases agree?”

Until now, it has been sufficient to use simple superscripted reference points such as ρ^f and ρ^g . However, in crossing dimensions in the order f, g, f , we probably should not assume that ρ^g is the same when comparing going from f to g with going from g to f , and the data agree with this possibility. We distinguish these as, respectively, $\rho^{f,g}$ and $\rho^{g,f}$.

Proposition 4. Assuming that the general model (Representation 1) holds, then

$$x_{p,q}^{f,f,f} = x_{p,q}^{f,g,f},$$

if and only if

$$\rho^{f,g} = \rho^{g,f}.$$

Note that the Narens (1996) theory has no reference points, and so it predicts that this agreement must hold. However, the data do not support that prediction. The above argument is depicted in Figure 3.

Method

The goal of the experiment is to test the same scale hypothesis of the Cross-Frequency Commutativity section as well as whether two reference points on different frequencies agree or not (see the Comparison of $f \rightarrow f \rightarrow f$ with $f \rightarrow g \rightarrow f$ section).

Respondents

A total of 12 respondents—mostly graduate and undergraduates from the University of California, Irvine—participated in the experiment reported. All respondents reported normal hearing. For logistical reasons, not every respondent participated in every condition. All respondents, except the participating coauthor (R22), received compensation of \$10 per session. Each person provided written consent and was treated in accordance with the Ethical Principles of Psychologists and Code of Conduct (American Psychological Association, 2002). Consent forms and procedures were approved by the University of California, Irvine’s Institutional Review Board. Respondents are given unique identifiers of the

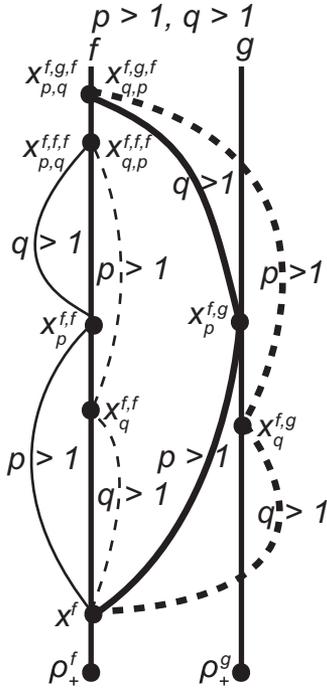


Figure 3. Depicted is the condition of Proposition 3. The narrow lines (solid and dotted) are on a single frequency, and the thicker lines (solid and dotted) go across the two frequencies, f, g . Independent of thickness, the solid lines depict the case of magnitude production with p followed by q , whereas the dotted line depicts magnitude production starting with q followed by p . Commutativity in the single frequency case holds if $x_{p,q} \sim x_{q,p}$, for the cross frequency case if $x_{p,q}^{f,g,f} \sim x_{q,p}^{f,g,f}$, and $\rho^{f,g} \sim \rho^{g,f}$ holds only if $x_{p,q} \sim x_{q,p} \sim x_{p,q}^{f,g,f} \sim x_{q,p}^{f,g,f}$.

form Rm , where m is a counter signifying only when an individual came to participate in any experiment we have conducted.

Stimuli

The basic signal was a sinusoidal tone presented for 100 ms, which included 10-ms on and off ramps. Because the theory is cast in terms of intensity increments above threshold intensity, for a left ear with a threshold of x_τ and a right ear threshold of u_τ , the effective stimulus (x, u) consists of $x = x' - x_\tau$ and $u = u' - u_\tau$, where (x', u') are the actual intensities presented. However, because all signals were well above threshold and the respondents were selected for normal hearing, the error in reporting intensities (x', u') in dB SPL (henceforth dB) is negligible. The basic stimulus consisted of two signals (tones), (x, x) , i.e., the same tone in both ears, and (z, z) , separated by 450 ms (see the Procedure section for additional details).

Apparatus

Stimuli were generated digitally using a personal computer and were played through a 24-bit digital-to-analog converter (RP2.1 real-time processor, Tucker-Davis Technology). Intensity and frequency were controlled through a programmable interface for the RP2.1, and stimuli were presented over Sennheiser HD265L headphones to the respondent seated in an individual, single-walled

IAC sound booth located in a quiet lab room. A safety ceiling of 90 dB was imposed in all experiments.

Statistical Method and Presentation of Results

We have no a priori model of how the data from individuals relate to one another, therefore all data analysis is done on individual data (e.g., Luce, 1995, p. 20).

We seek to evaluate parameter-free null hypotheses that have the generic form $L_{\text{side}} = R_{\text{side}}$. As a matter of logic, a null hypothesis can never be proven empirically, but it can be empirically supported. At present, we have no model of error for respondent-produced estimates. This precludes any principled use of parametric testing. Further, as far as we are aware, but without intending any comment on the choices made by other authors, there is presently no clear or agreed upon approach to this issue in the psychological testing of representational models. The problem is common in physics, and the solution in that field takes the form of providing a criterion by which data can be said to support (or not support) a tested null hypothesis. To this end, we developed a three-pronged criterion.

Statistical method. Lack of parametric information leads us to the use of a nonparametric test. Our choice is the Mann-Whitney U at a significance level of .05. This was the choice of others in similar studies (e.g., Augustin & Maier, 2008; Ellermeier & Faulhammer, 2000; Ellermeier, Narens, & Dielmann, 2003; Steingrimsson, 2009, in press; Steingrimsson & Luce, 2005a, 2005b, 2006, 2007; Zimmer, 2005; Zimmer, Luce, & Ellermeier, 2001). Because respondents adjust intensities in discrete steps and whose estimates appear reasonably Gaussian, medians are known to be well estimated by the mean, and thus the central tendencies are reported as means with variability in standard deviations—the test itself, being a rank-order one, is on the medians. Put more formally, the tested version of Representation 1 is

$$W(p) = \frac{\Psi(\bar{x}_p) - \Psi(\rho_i)}{\Psi(x) - \Psi(\rho_i)}, \quad (9)$$

where \bar{x}_p is the median produced value.

Sample size adequacy. To evaluate whether the sample size is sufficiently large to detect a true failure of the null hypothesis, all statistical results were verified using Monte Carlo simulations (see Steingrimsson, 2009, and Steingrimsson & Luce, 2005a for details).

Effect size. Currently, no accepted method for calculating effect size exists for the Mann-Whitney U. Instead, we use the simple observation that if two medians (means) differ by less than Weber's fraction, they are arguably not noticeably different to an observer.² Teghtsoonian (1971) reported that Weber's fractions for loudness from three (reportedly conservative and independent) studies all agreed on approximately .05.

Our criterion for saying that a result supports our hypothesis is for all three indicators to agree; otherwise, it is not supported.

Procedure

Experiments were conducted in sessions of at most one hour each. The initial session was devoted to obtaining written consent, explain-

² We thank J. Yellott for this simple and elegant observation.

ing the task, and running practice trials. All respondents trained for one additional session. Rest periods were encouraged but both their frequency and duration were under the respondent's control.

With reference to Figure 2, Panel 1, testing the same scale hypothesis consists of presenting a standard x^f and, following the solid line, obtaining first an estimate of $x_p^{f,g}$, then using $x_p^{f,g}$ as a standard, obtaining an estimate of $x_{p,q}^{f,g,g}$. Following the dotted line, the corresponding estimates are $x_q^{f,g}$ then $x_{q,p}^{f,g,g}$. The statistical hypothesis is $x_{p,q}^{f,g,g} \approx x_{q,p}^{f,g,g}$.

We use a two-ear presentation and so the standard is (x^f, x^f) , which means the joint presentation of the tone x^f in the left ear and tone x^f in the right ear. Let $\langle (x^f, x^f), (z^g, z^g) \rangle$ mean that the joint presentation (x^f, x^f) is played, followed by 450 ms of the joint presentation (z^g, z^g) —a tone is a sinusoid of 100-ms duration, including 10-ms on and off ramps (see the Stimulus section). The intensity of the tone z^g is under the respondent's control, such that s/he can either increase or decrease its intensity. The intensity changes are in step sizes of 0.5, 1, 2, or 4 dB, each tied to the keys A, S, D, and F for increasing intensity and the semicolon, L, K, and J keys for decreasing intensity. After each key press, the altered tone sequence was replayed. In addition, respondents could replay the tone sequence unaltered by pressing the R key.

To produce the estimate for $x_p^{f,g}$, the respondents adjusted the intensity of z^g until the stimulus (z^g, z^g) sounded some prescribed proportion p of the standard stimulus (x^f, x^f) . The respondent could adjust z^g as often as desired, but when satisfied with the production, s/he would press the B key. The final value of z^g was taken as an estimate for $x_p^{f,g}$. In a following step, this estimate was used as a standard $(x_p^{f,g}, x_p^{f,g})$, using the same process to obtain an estimate of $x_{p,q}^{f,g,g}$. The respondent produced estimates for $x_q^{f,g}$ and $x_{q,p}^{f,g,g}$ in an analogous fashion. Thus, arriving at an estimate for $x_{p,q}^{f,g,g}$ and $x_{q,p}^{f,g,g}$ required a total of four individual estimates.

In most cases, two instantiations of the standard, (x^f, x^f) and (f^f, f^f) , were used and randomly interleaved within a block of trials that thus consisted of eight estimates (four for each standard) to arrive at two tests of the statistical hypothesis, namely, $x_{p,q}^{f,g,g} \approx x_{q,p}^{f,g,g}$ and $f_{p,q}^{f,g,g} \approx f_{q,p}^{f,g,g}$, respectively.

To avoid intersession variability, the eight estimates were collected within a block of trials (See Appendices A.1 and A.4 in Steingrímsson & Luce, 2005a, for details). Respondents typically completed between eight and 10 blocks per session. The value of the current proportion was displayed on the monitor, in addition to the task instructions and the current block number and trial number. The typical number of estimates for each of $x_{p,q}^{f,g,g}$ and $x_{q,p}^{f,g,g}$ and of $f_{p,q}^{f,g,g}$ and $f_{q,p}^{f,g,g}$ were around 30 (excluding training), requiring three to four sessions to collect data for each condition.

Design and Condition

The testing strategy aimed at evaluating the same scale hypothesis in three ways:

1. Test the three p, q relationships of Proposition 3.

Case 1. $p \geq 1, q \geq 1$ (see Figure 2, Panel 1).

Case 2. $p' < 1, q' < 1$ (see Figure 2, Panel 2).

Case 3. $p \geq 1 > q'$ (see Figure 2, Panel 3).

2. Test of Proposition 4, whether $\rho^{f,g} = \rho^{g,f}$, i.e., is the reference point the same on any two frequencies (see Figure 3).
3. Test the conditions for several frequencies and a large frequency range.

Listed in Table 1, are the 13 stimulus conditions that together address these three aims.

Aim 3 is addressed in various ways: Magnitude productions are made both from higher to lower frequencies (Conditions 1, 2, 3, 4, 6, 7, 8, 19, 12, 13) and lower to higher frequencies (Conditions 3, 4, 9, 11) and over a wide range of frequencies (Conditions 6, 7). For all except Conditions 12 and 13, two standards are used for each condition, resulting in two tests for each condition.

Results

The results are summarized in Table 2. The detailed results are listed in Appendix B, in Tables B1 and B2. The condensed presentation here avoids burdening the reader with a list of 68 different tests, where as Tables B1 and B2 ensure full reporting of the appropriate details.

The pattern of results in Table 2 is clear: The same scale hypothesis of Proposition 3 is well supported for Cases 1 and 2, whereas Case 3 is unambiguously rejected (corresponding to Conditions 1, 2, and 3 of Figure 2). The failure of Case 3 is evidence against the hypothesis that $\rho_+ = \rho_-$ for both frequencies. For the test of Proposition 4 (corresponding to Figure 3), the data do not support the general hypothesis that $\rho^{f,g} = \rho^{g,f}$.

From Tables B1 and B2, it can be seen that the two of the failures in test of Case 1 (see Table 2) are for the same respondent in the same condition (failure for each of the two standards). Otherwise, there does not seem to be any pattern to the failures. Nine respondents attempted to exceed the 90-dB safety limit. For three, the standard was lowered, and the problem resolved itself. Three participated only in a test of Case 2, in which the proportions were less than one, in which case the safety limit is not a limiting factor; the other respondents were unavailable to provide additional data. One respondent appeared not to have understood the instructions and showed dramatic intersession variability (~ 10 dB) in three sessions. Those data are not reported.

Discussion and Conclusions

1. We began by showing how the ratio scale, commutativity property of prosthetic attributes for signals that vary only in intensity can be extended to signals that vary both in intensity and in another variable, such as frequency. Those results were the basis of the experimental program.
2. Empirical evidence supports the notion that for $p, q \geq 1$ and $p', q' < 1$, individuals rely on a single scale for loudness regardless of stimulus frequency.
3. The evidence is consistent with the idea that reference points for $p' < 1$ and for $p \geq 1$ do differ.
4. The evidence is consistent with reference points for different frequencies do differ. Specifically, $\rho^f \neq \rho^g$, and $\rho^{f,g} \neq \rho^{g,f}$.

Table 1
The Experimental Condition Under Which the Hypotheses of Propositions 3 and 4 Were Tested

Condition	Proportion		Frequency (Hz)		Standard (dB)	
	p	q	f	g	x^f	t^f
Proposition 3, Case 1: $p > 1, q > 1$						
1 ^a	2	3	1,000	2,000	54	50
2	2	3	1,000	2,000	54	50
3	2	3	1,000	2,000 (x)/500 (u)	54	50
4	150%	200%	1,000	2,000	54	50
5	150%	200%	1,000	500	54	50
6	150%	200%	200	5,000	54	50
Proposition 3, Case 2: $p' < 1, q' < 1$						
7	75%	50%	200	5,000	82	72
8	75%	50%	1,000	2,000	82	76
9	75%	50%	1,000	500	82	72
Proposition 3, Case 3: $p' < 1, q > 1$						
10	75%	150%	1,000	2,000	70 (60) ^b	74 (64) ^b
11	75%	150%	1,000	500	70 (60) ^b	74 (64) ^b
Proposition 4: Equality of reference points, $\rho^{f,g} = \rho^{g,f}$						
12	150%	200%	1,000	2,000	50	none
13	50%	75%	1,000	2,000	78	none

Note. First listed is a numerical identifier for each testing condition, followed by the values of the proportions, followed by the frequencies and, finally, the intensity of the two standards.

^a Presented in left ear only. ^b The intensity in parenthesis is the one used for a single respondent.

- We asked, "Is loudness an intensity scale that is independent of frequency?" The data suggested the answer is yes as long as reference points are not assumed to remain fixed independent of procedure.

Several directions need to be explored. One that has so far escaped us is a principled theory of reference points. They are clearly important in our data, and we have only been able to treat them as parameters to be estimated. A second question is, "Do the same results hold for other prothetic continua, among them brightness of monochromatic colors?" If so, researchers should expand the exploration to other prothetic continua that have been shown to be systematic using Stevens' magnitude methods. For those to which it does apply, the next step is to extend the model to intermodal situations. Should that work out, the implication is that there is a single notion of subjective intensity for a person, a somewhat sweeping idea.

Perceptual independence and perceptual separability have been studied for many years by many investigators (see Ashby &

Townsend, 1986, for a review). This literature focuses on the processing of perceptual information and the modeling of relationships among perceptual dimensions, perceptual processing, decision processing, and the types of responses. The idea is to use concepts and results from information theory and statistical modeling to understand the relative amounts independence and redundancy (sometimes called separable and integrable) of physical dimensions at various stages of perceptual and cognitive processing. In contrast, our modeling procedures are derived from a measurement theory developed by Luce and Narens (Luce & Narens, 1985; Narens & Luce, 1976) that generalizes the measurement and interactions of physical dimensions in a way designed to have possible applications to the behavioral sciences. This is accomplished in a manner that, as in physics, is completely independent of statistical and information theory. Our use of reference points may be particularly relevant to theories of perceptual independence and perceptual separability. At this time, we do not have

Table 2
Summary of the Testing of the Commutativity Hypotheses of Propositions 3 and 4

Test of	No. of tests	No. hold	No. fail	% hold	Hypothesis
Proposition 3, Case 1: $p > 1, q > 1$	38 ^a	34	4	89	Supported
Proposition 3, Case 2: $p' < 1, q' < 1$	12 ^a	11	1	91	Supported
Proposition 3, Case 3: $p' < 1, q > 1$	8	0	8	0	Rejected
Proposition 4: Equality of reference points, $\rho^{f,g} = \rho^{g,f}$	8	1	7	14	Rejected

Note. First listed is the specific test conducted, followed by the number of tests of each, how many of those tests were found to hold, how many failed, and the percentage of tests that held of the total number of tests. Finally, the conclusion about each hypothesis is listed.

^a Cases 1 and 2 include data from C_{12} and C_{13} that primarily test Proposition 4.

a good measurement-theoretic understanding of the reference points, other than that they are parameters that allow the commutativity hypothesis to account for a far broader set of results.

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(Appendices follow)

Appendix A

Proofs

Proposition 1

Case 1

This is the case $x \rightarrow x_p$ followed by $x_p \rightarrow x_{p,q}$:

$$W(p) = \frac{\psi(x_p) - \psi(\rho_+)}{\psi(x) - \psi(\rho_+)},$$

$$W(q) = \frac{\psi(x_{p,q}) - \psi(\rho_+)}{\psi(x_p) - \psi(\rho_+)}.$$

So,

$$\begin{aligned} W(p)W(q) &= \left(\frac{\psi(x_p) - \psi(\rho_+)}{\psi(x) - \psi(\rho_+)} \right) \left(\frac{\psi(x_{p,q}) - \psi(\rho_+)}{\psi(x_p) - \psi(\rho_+)} \right) \\ &= \frac{\psi(x_{p,q}) - \psi(\rho_+)}{\psi(x) - \psi(\rho_+)}. \end{aligned}$$

Similarly, for the order q, p ,

$$\begin{aligned} W(q)W(p) &= \left(\frac{\psi(x_q) - \psi(\rho_+)}{\psi(x) - \psi(\rho_+)} \right) \left(\frac{\psi(x_{q,p}) - \psi(\rho_+)}{\psi(x_p) - \psi(\rho_+)} \right) \\ &= \frac{\psi(x_{q,p}) - \psi(\rho_+)}{\psi(x) - \psi(\rho_+)}. \end{aligned}$$

By the commutativity of multiplication, equate these, and it is immediate that $x_{p,q} = x_{q,p}$, i.e., commutativity holds.

Case 2

Here we have $p' < 1$, $q' < 1$, and a completely analogous argument using ρ_- again forces commutativity.

Case 3

Now assume we are working with $p > 1 > q'$. Then, according to the model, when the presentation order is p, q' , we have

$$W(p) = \frac{\psi(x_p) - \psi(\rho_+)}{\psi(x) - \psi(\rho_+)},$$

$$W(q') = \frac{\psi(x_{p,q'}) - \psi(\rho_-)}{\psi(x_p) - \psi(\rho_-)}.$$

Therefore

$$W(p)W(q') = \left(\frac{\psi(x_p) - \psi(\rho_+)}{\psi(x) - \psi(\rho_+)} \right) \left(\frac{\psi(x_{p,q'}) - \psi(\rho_-)}{\psi(x_p) - \psi(\rho_-)} \right). \quad (\text{A1})$$

Next, consider the presentation order q', p :

$$\begin{aligned} W(q') &= \frac{\psi(x_{q'}) - \psi(\rho_-)}{\psi(x) - \psi(\rho_-)}, \\ &= \frac{\psi(x_{q',p}) - \psi(\rho_+)}{\psi(x_{q'}) - \psi(\rho_+)}. \end{aligned}$$

Thus,

$$W(q')W(p) = \left(\frac{\psi(x_{q'}) - \psi(\rho_-)}{\psi(x) - \psi(\rho_-)} \right) \left(\frac{\psi(x_{q',p}) - \psi(\rho_+)}{\psi(x_{q'}) - \psi(\rho_+)} \right). \quad (\text{A2})$$

Equating Equation A1 to Equation A2, $x_{p,q'} = x_{q',p}$ holds iff $\rho_+ = \rho_-$, as claimed.

Proposition 2

By the proof of Proposition 1, we have for the two p, q orders,

$$\begin{aligned} \tau_+(p, q) &= W(p)W(q) = \frac{\psi(x_{p,q}) - \psi(\rho_+)}{\psi(x) - \psi(\rho_+)} \\ \tau_+(q, p) &= W(q)W(p) = \frac{\psi(x_{q,p}) - \psi(\rho_+)}{\psi(x) - \psi(\rho_+)}. \end{aligned} \quad (\text{A3})$$

Because we know that $x_{p,q} = x_{q,p}$, it suffices to look only at the first display. Holding p and q fixed, we may vary x over several values and determine $x_{p,q}(x)$ empirically. Thus, we have a set of equations with two parameters, $\tau_+(x)$ and ρ_+ . Assume Equation 4 (i.e., ψ is a power function, and so Equation A3 can be rewritten

$$\rho_+^\beta = \frac{x^\beta \tau_+(p, q) - x_{p,q}^\beta}{\tau_+(p, q) - 1}, \quad (\text{A4})$$

which permits us to estimate ρ_+ . Exactly analogous calculations hold for $p' < 1$, $q' < 1$, with parameters $\tau_-(p', q')$ and ρ_- .

For $p' < 1$ and $q' < 1$, the monotonicity of W implies

$$\tau_-(p', q') < \tau_+(p, q).$$

Proposition 3

Case 1

For the mapping $f \rightarrow g \rightarrow g$ (i.e., where $x^f \rightarrow x_p^{f,g}$ and $x_p^{f,g} \rightarrow x_{p,q}^{f,g,g}$), we have

$$\begin{aligned} W(p) &= \frac{\psi(x_p^{f,g}) - \psi(\rho_+^{f,g})}{\psi(x^f) - \psi(\rho_+^f)}, \\ &= \frac{\psi(x_{p,q}^{f,g,g}) - \psi(\rho_+^{g,g})}{\psi(x_p^{f,g}) - \psi(\rho_+^{f,g})}. \end{aligned}$$

(Appendices continue)

Thus,

$$\begin{aligned} W(p)W(q) &= \left(\frac{\psi(x_p^{f,g}) - \psi(\rho_+^{f,g})}{\psi(x^f) - \psi(\rho_+^f)} \right) \left(\frac{\psi(x_{p,q}^{f,g,g}) - \psi(\rho_+^{g,g})}{\psi(x_p^{f,g}) - \psi(\rho_+^{f,g})} \right) \\ &= \frac{\psi(x_{p,q}^{f,g,g}) - \psi(\rho_+^{g,g})}{\psi(x^f) - \psi(\rho_+^f)}. \end{aligned}$$

Similarly, for the order q, p ,

$$\begin{aligned} W(q)W(p) &= \left(\frac{\psi(x_{q,p}^{f,g,g}) - \psi(\rho_+^{g,g})}{\psi(x_q^{f,g}) - \psi(\rho_+^{f,g})} \right) \left(\frac{\psi(x_q^{f,g}) - \psi(\rho_+^{f,g})}{\psi(x^f) - \psi(\rho_+^f)} \right) \\ &= \frac{\psi(x_{q,p}^{f,g,g}) - \psi(\rho_+^{g,g})}{\psi(x^f) - \psi(\rho_+^f)}. \end{aligned}$$

By the commutativity of multiplication, equate these, and it is immediate that

$$x_{p,q}^{f,g,g} = x_{q,p}^{f,g,g},$$

i.e., commutativity holds.

Cases 2 and 3

Case 2 is analogous. And following the pattern of Proposition 1, Case 3 requires $\rho_+ = \rho_-$ for both frequencies.

Proposition 4

For the case $f \rightarrow g \rightarrow f$ (i.e., where $x^f \rightarrow x_p^{f,g}$ and $x_p^{f,g} \rightarrow x_{p,q}^{f,g,f}$), we have for $i = +, -$

$$W(p) = \frac{\psi(x_p^{f,g}) - \psi(\rho_i^{f,g})}{\psi(x^f) - \psi(\rho_i^f)},$$

$$W(q) = \frac{\psi(x_{p,q}^{f,g,f}) - \psi(\rho_i^{g,f})}{\psi(x_p^{f,g}) - \psi(\rho_i^{f,g})}.$$

So

$$\begin{aligned} W(p)W(q) &= \left(\frac{\psi(x_p^{f,g}) - \psi(\rho_i^{f,g})}{\psi(x^f) - \psi(\rho_i^f)} \right) \left(\frac{\psi(x_{p,q}^{f,g,f}) - \psi(\rho_i^{g,f})}{\psi(x_p^{f,g}) - \psi(\rho_i^{f,g})} \right) \\ &= \left(\frac{\psi(x_{p,q}^{f,g,f}) - \psi(\rho_i^{g,f})}{\psi(x^f) - \psi(\rho_i^f)} \right) \left(\frac{\psi(x_p^{f,g}) - \psi(\rho_i^{f,g})}{\psi(x_p^{f,g}) - \psi(\rho_i^{f,g})} \right). \end{aligned}$$

For the case $f \rightarrow f \rightarrow f$ (Case 1 of Proposition 1), yielded

$$W(p)W(q) = \frac{\psi(x_{p,q}^{f,f,f}) - \psi(\rho_i^f)}{\psi(x^f) - \psi(\rho_i^f)}$$

which means

$$x_{p,q}^{f,g,f} \approx x_{p,q}^{f,f,f} \text{ iff } \rho_i^{f,g} \approx \rho_i^{g,f}.$$

(Appendices continue)

Appendix B

Detailed Results

Tables B1 and B2 contain the detailed results of tests conducted.

Table B1
Tests of Proposition 3

Condition	Respondent	Sound level (dB)								n	p_x	p_t
		$x_{p,q}^{f.g.g}$		$x_{q,p}^{f.g.g}$		$t_{p,q}^{f.g.g}$		$t_{q,p}^{f.g.g}$				
		M	SD	M	SD	M	SD	M	SD			
Proposition 3, Case 1: $p > 1, q > 1$												
C ₁	10	85.43	3.65	85.37	3.59	84.60	3.18	84.25	4.03	30	.847	.929
	22	78.10	5.47	78.53	4.78	77.22	5.69	76.10	5.00	30	.784	.374
	75	71.33	3.49	71.78	3.41	70.75	4.39	71.80	3.86	30	.711	.205
C ₂	56	78.40	5.89	80.45	5.05	78.93	6.92	79.12	6.00	30	.151	.959
	66	82.07	4.18	81.13	5.40	79.40	2.86	79.53	3.75	30	.542	.726
	71	86.77	3.33	87.63	1.96	86.97	1.76	85.85	3.18	30	.372	.188
C ₃	75	75.88	4.36	76.16	3.62	74.86	3.30	76.09	4.99	37	.618	.137
	66	77.85	5.41	78.19	2.99	82.00	2.53	82.46	4.46	30	.646	.544
	75	72.47	3.69	70.98	4.68	73.63	4.24	74.30	4.57	30	.396	.311
C ₄	47	78.23	4.89	77.93	5.06	76.60	4.25	77.13	5.51	30	.635	.721
	66	80.85	4.14	81.30	5.02	79.07	4.90	79.23	4.40	28	.941	.875
	68	77.83	5.90	75.55	6.61	74.85	6.52	73.55	4.59	30	.157	.148
C ₅	71	86.47	2.35	86.17	2.27	86.37	3.06	84.98	2.48	30	.460	.021*
	75	78.53	5.82	78.12	5.31	77.93	5.60	78.52	6.03	28	.761	.824
	76	86.63	2.00	85.20	2.77	86.41	1.82	84.39	5.17	27	.075	.834
C ₆	75	80.78	2.94	81.89	3.96	79.09	3.96	81.84	3.89	32	.107	.017*
	76	79.52	3.54	82.14	3.00	79.25	4.17	81.70	3.56	28	.001**	.012*
C ₆	10	86.53	5.81	86.89	4.38	83.68	5.41	84.07	4.70	28	.497	.863
Proposition 3, Case 2: $p' < 1, q' < 1$												
C ₇	47	71.30	6.20	71.40	3.56	57.80	8.61	57.30	8.57	30	.946	.894
C ₈	59	75.92	4.81	74.23	4.24	72.93	7.00	70.50	5.69	30	.123	.082
	76	52.40	7.48	48.75	10.20	48.84	12.45	48.28	7.42	32	.081	.572
C ₉	59	75.87	3.45	72.62	5.41	70.48	6.54	69.87	7.63	30	.023*	.976
	63	50.83	10.29	50.08	13.10	47.65	10.49	47.58	12.71	30	.496	.953
Proposition 3, Case 3: $p' < 1, q > 1$												
C ₁₀	59	77.83	2.66	87.70	3.27	77.00	2.48	87.77	2.95	30	<.001**	<.001**
	76	64.35	3.41	75.25	4.11	62.30	3.88	76.05	2.74	30	<.001**	<.001**
C ₁₁	59	75.20	5.03	83.50	3.94	75.30	4.09	84.70	2.57	30	<.001**	<.001**
	76	63.65	3.73	76.27	4.00	61.36	4.60	77.10	3.59	29	<.001**	<.001**

Note. The table lists by condition and by respondent the detailed results of the testing. For each condition, the columns are as follows: the participating respondent, his/her averaged estimates, and their standard deviations, as well as the number of observation produced by each person. The final two columns list the results of the statistical testing where the notation p_x means $x_{p,q}^{f.g.g} = x_{q,p}^{f.g.g}$, and p_t means $t_{p,q}^{f.g.g} = t_{q,p}^{f.g.g}$.

* Rejection at the .05 limit. ** Rejection at the .01 limit.

(Appendices continue)

Table B2

Tests of Proposition 4: Equality of Reference Points $\rho_{f,g} \approx \rho_{g,f}$

Condition	Respondent	Frequency	Sound level (dB)				n	p_{\rightarrow}	p_{\times}
			$x_{p,q}$		$x_{q,p}$				
			M	SD	M	SD			
Proposition 4, Case 1: $p > 1, q > 1$									
C ₁₂	10	$f \rightarrow f \rightarrow f$	80.02	2.43	79.17	3.37	34	.279	.005
		$f \rightarrow g \rightarrow f$	83.06	4.00	83.47	3.77		.922	<.001
	22	$f \rightarrow f \rightarrow f$	71.70	3.58	72.30	3.12	30	.563	.001
		$f \rightarrow g \rightarrow f$	75.77	4.06	75.25	4.22		.733	.005
Proposition 4, Case 2: $p' < 1, q' < 1$									
C ₁₃	76	$f \rightarrow f \rightarrow f$	35.11	4.50	35.71	3.55	28	.873	<.001
		$f \rightarrow g \rightarrow f$	40.68	4.99	39.96	3.19		.423	<.001
	22	$f \rightarrow f \rightarrow f$	65.52	3.31	64.64	2.04	28	.702	.286
		$f \rightarrow g \rightarrow f$	67.39	3.01	66.44	3.22		.120	.007

Note. The table lists by condition and by respondent the detailed results of the testing. For each condition, the columns are as follows: the participating respondent, his/her averaged estimates, and their standard deviations, as well as the number of observation produced by each person. The final two columns list the results of the statistical testing where the notation p_{\rightarrow} means either $x_{p,q}^{f,f,f} = x_{q,p}^{f,f,f}$ or $x_{p,q}^{f,g,f} = x_{q,p}^{f,g,f}$, and p_{\times} means either $x_{p,q}^{f,f,f} = x_{q,p}^{f,g,f}$ or $x_{p,q}^{f,g,f} = x_{q,p}^{f,f,f}$.

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